

RESEARCH ARTICLE

Simulation Study of the Test on Covariance Estimator for Outlier Detection in Multivariate Data with Mean and Covariance Shifts

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Abstract Outlier detection in multivariate data is complex compared to univariate data, which can be done using graphical inspection. Outlier detection is also one of the common issues in multivariate analysis and has been applied to tax fraud detection and industrial food inspection. Outliers' studies are closely related to robust estimators of the sample mean and covariance matrix as these estimators are resistant toward outliers. The Test on Covariance (TOC) is a newly developed robust estimator for multivariate data. Until now, TOC's performance was investigated for two outlier scenarios by shifting the mean and covariance separately. TOC shows good results in both outlier scenarios and is found to be applicable in detecting outliers. In this study, the performance of TOC is investigated further in detecting outliers via simulation study for other outliers' scenarios by shifting the mean and covariance simultaneously. Other robust estimators; Fast Minimum Covariance Determinant (FMCD), Minimum Vector Variance (MVV), Covariance Matrix Equality (CME) and Index Set Equality (ISE) are used as a comparison. Various conditions of sample sizes, n = 30, 50, 100, number of variables, p = 2, 3, 5 and percentage of outliers,

 $\varepsilon=5\%,15\%$ are considered in the simulation study. The performance of all robust estimators is measured by probability to detect outliers *(pout)*, masking error *(pmask)* and swamping error *(pswamp)*. Results present that the TOC can be the best robust estimator, give the same performance as other robust estimators in detecting outliers, and have a low masking error when outliers and inliers are far from each other. Moreover, TOC displays good results in low swamping errors for most cases which means TOC has a low probability of misclassifying inliers as outliers compared to other robust estimators. In conclusion, TOC is an applicable and promising approach for outlier detection in multivariate data and can be incorporated with other multivariate analyses.

Keywords: Test on Covariance, robust estimator, Mahalanobis distance, outliers, multivariate data.

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Introduction

Outliers are abnormal data points that deviate significantly from most of the observations in the dataset. While detecting outliers in univariate data is relatively straightforward and can often be done visually [1,2], identifying outliers in multivariate data is more complex. Outlier detection is also one of the common issues in multivariate analysis and has been applied to industrial food inspection [3] and tax fraud detection [4]. One standard method for detecting multivariate outliers is the Mahalanobis Distance (MD), which measures the distance of an observation from the data center while accounting for the data's overall shape [5]. The formula for MD is provided in equation (1),

$$d_{i}(\overline{\mathbf{x}}, \mathbf{S}) = \sqrt{(\mathbf{x}_{i} - \overline{\mathbf{x}})' \mathbf{S}^{-1}(\mathbf{x}_{i} - \overline{\mathbf{x}})}, \quad i = 1, 2,, n,$$
(1)

where $\overline{\mathbf{x}}$ is the sample mean, \mathbf{S} is the covariance matrix and n is the sample size. However, MD's sample mean and covariance matrix have masking and swamping effects when the multivariate data



contain outliers. Many studies have proposed using robust estimators of mean and covariance matrices as the robust estimators are resistant to outliers to solve masking and swamping problems [6,7].

Various robust estimators such as S, M, Method of Moments (MM), Minimum Volume Ellipsoid (MVE), Minimum Covariance Determinant (MCD) and Fast MCD (FMCD) estimators have been presented in the previous studies. Among these robust estimators, FMCD has been and still is widely used as shown in Salleh [8] and Mashuri *et al.* [9]. The Fast Minimum Covariance Determinant (FMCD) proposed by Rousseeuw & Van Driessen [10] is commonly used as FMCD has proved computationally efficient. The performance of FMCD uses clustered and shifted outliers and the combination of these outliers. However, FMCD still needs a lot of calculation and is time-consuming as FMCD uses covariance determinants in the last step of the algorithm [11].

Hence, Herwindiati *et al.* [6] took the initiative to propose robust estimators by using vector variance and called the new robust estimator Minimum Vector Variance (MVV). Regarding computation time, MVV is faster than FMCD but still lacking when the dimension or number of variables, *p* increases [12]. MVV has been tested to detect outliers for mean shift outlier scenarios only.

In 2013, Salleh [8] proposed two new robust estimators called Covariance Matrix Equality (CME) and Index Set Equality (ISE). The CME involves a comparison of two covariance matrices, element by element. Conversely, ISE is only a logical comparison of two index sets: old subset and new subset. ISE has been demonstrated to work excellently in computation time [11]. However, Salleh [8] did not test CME and ISE to detect outliers in multivariate data but applied CME and ISE to monitor process variability. According to Salleh [8], finding a condition for two covariance matrices to be equal can be further examined.

Hence, Abd Mutalib *et al.* [13] proposed a new robust estimator based on the idea of CME and ISE named Test on Covariance (TOC) and the performance of TOC is investigated via a simulation study. Abd Mutalib *et al.* [13] used one outlier scenario named mean shift in their study. Next, Abd Mutalib *et al.* [14] study the performance of TOC in two outlier scenarios: mean shift and covariance shift separately. Both studies found that TOC shows good results and a promising approach to detecting outliers for multivariate data. Abd Mutalib *et al.* [15] conducted a further study to investigate the performance of TOC in five real multivariate datasets from existing literature. Findings showed that TOC is a promising approach to detecting outliers in all datasets.

Therefore, in this study, further investigation is conducted to discover the performance of TOC in outlier scenarios when both the mean and covariance are shifted simultaneously. The performance of TOC will be analyzed and compared with FMCD, MVV, CME and ISE. The rest of the paper is organized as follows. The following section explains the materials and methods used in this study. Details about the simulation study and TOC are discussed in detail in this section. Then, the results and discussion of the simulation study are presented next. The last section presents the conclusion of this study.

Materials and Methods

Figure 1 shows the general procedure of this study's simulation study. Multivariate data is generated first according to the outlier scenario. Next, TOC and existing robust estimators are obtained. The simulation study is done 10,000 times after obtaining the robust estimators. Lastly, the performance of TOC is evaluated and compared with other robust estimators using performance measurement by Sebert *et al.* [16].

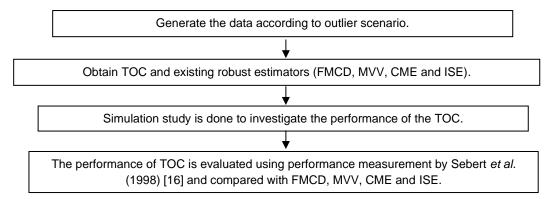


Figure 1. Simulation study general procedure



Test on Covariance Estimator

The TOC involving the equality test of variance-covariance structure is proposed by Abd Mutalib *et al.* [13]. The equality of two covariance structures is tested by using equation (2) with the hypothesis $H_0: \Sigma_{old} = \Sigma_{new}$ versus $H_1: \Sigma_{old} \neq \Sigma_{new}$,

$$u = \upsilon \left[\sum_{i=1}^{p} (\lambda_i - \ln \lambda_i) - p \right], \tag{2}$$

where $\upsilon=n-1$ is the degrees of freedom, p is the number of variable and λ_i are the eigenvalue of $\Sigma_{new}\Sigma_{old}^{-1}$. Σ_{old} and Σ_{new} are obtained from H_{old} and H_{new} from the algorithm. The null hypothesis, H_0 is rejected if $u>\chi^2\left[\alpha,\frac{1}{2}\,p\,\big(\,p+1\big)\right]$ [17].

The TOC algorithm is similar to the FMCD algorithm, with a new procedure introduced for the stopping rule in Step 6. The FMCD algorithm is outlined below.

Step 1: Select an arbitrary subset H_{old} containing h different observations, where h is the smallest integer greater than or equal (n+p+1)/2, where p is the number of variables and n is the sample size.

Step 2: Compute the mean vector $\bar{X}_{H_{old}}$ and covariance matrix $S_{H_{old}}$ of all observations belonging to H_{old} .

 $\textbf{Step 3: Compute} \ \, d_{H_{old}}^2 \left(i \right) \! = \! \left(X_i - \overline{X}_{H_{old}} \right)' S_{H_{old}}^{-1} \left(X_i - \overline{X}_{H_{old}} \right) \ \, \text{for} \ \, i = 1, 2, \mathbf{K} \ \, , n \, .$

Step 4: Sort $d_{H_{old}}^2\left(i\right)$ for i=1,2,K, n in increasing order $d_{H_{old}}^2\left(\pi\left(1\right)\right) \leq d_{H_{old}}^2\left(\pi\left(2\right)\right) \leq K \leq d_{H_{old}}^2\left(\pi\left(n\right)\right)$ where π is a permutation on i=1,2,K, n.

Step 5: Define $H_{new} = \left\{ X_{\pi(1)}, X_{\pi(2)}, \mathbf{K}, X_{\pi(h)} \right\}$ and then calculate $\overline{X}_{H_{new}}, S_{H_{new}}$ and $d_{H_{new}}^2 \left(i \right)$ for $i = 1, 2, \mathbf{K}, n$.

 $\begin{array}{lll} \textbf{Step 6}_{\text{FMCD}} \colon & \text{Stopping Rule. If } \det\left(S_{H_{new}}\right) = 0, & \text{repeat Step 1} - \text{Step 5. Otherwise, if} \\ \det\left(S_{H_{new}}\right) < \det\left(S_{H_{old}}\right), & \text{let } H_{old} \coloneqq H_{new}, \\ \overline{X}_{H_{old}} \coloneqq \overline{X}_{H_{new}} & \text{and } S_{H_{old}} \coloneqq S_{H_{new}}. & \text{Then go to Step 3.} \\ \\ \text{Otherwise, the process is stopped and } \det\left(S_{H_{new}}\right) = \det\left(S_{H_{old}}\right) & \text{is obtained.} \\ \end{aligned}$

Step 6 for TOC is given as follows,

Step 6_{TOC} (Stopping Rule): If H_0 is rejected, calculate $\overline{X}_{H_{new}}$ and let $H_{old} \coloneqq H_{new}$, $\overline{X}_{H_{old}} \coloneqq \overline{X}_{H_{new}}$ and $S_{H_{old}} \coloneqq S_{H_{new}}$. Then go to Step 3. Otherwise, the process is stopped.

The complete step of the TOC algorithm is shown in Figure 2.



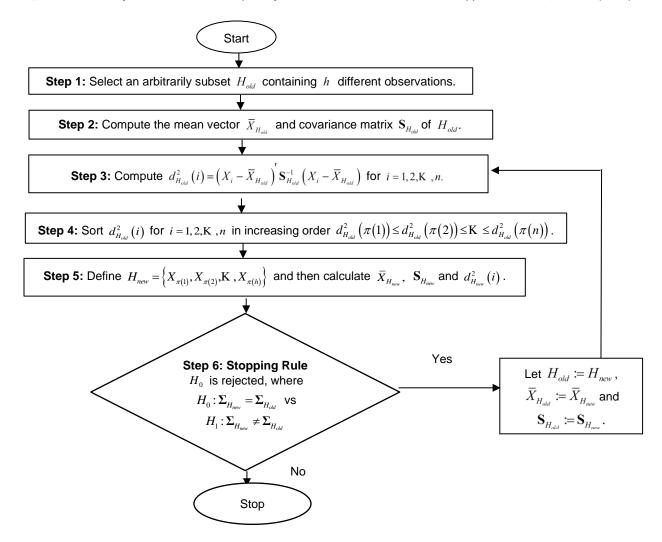


Figure 2. TOC Algorithm

In this study, the performance of TOC in detecting outliers in multivariate data is compared with FMCD, MVV, CME and ISE. The MVV, CME, and ISE also used the FMCD algorithm but differed in Step 6. The following are Step 6 for MVV, CME and ISE.

Step 6_{MVV}: If $Tr\left(S_{H_{new}}^2\right) = 0$, repeat Step 1 - Step 5. Otherwise, if $Tr\left(S_{H_{new}}^2\right) \neq Tr\left(S_{H_{old}}^2\right)$, let $H_{old} \coloneqq H_{new}$, $\overline{X}_{H_{old}} \coloneqq \overline{X}_{H_{new}}$ and $S_{H_{old}} \coloneqq S_{H_{new}}$. Then go to Step 3. Otherwise, the process is stopped and $Tr\left(S_{H_{new}}^2\right) = Tr\left(S_{H_{new}}^2\right)$ is obtained.

$$\text{Step 6cme: If } \sqrt{Tr \Big(S_{H_{new}} - S_{H_{old}}\Big)^2} \neq 0, \quad I_{new} \neq I_{old}$$

 $\textbf{Step 6}_{\text{ISE}} : \text{If } I_{\textit{new}} \neq I_{\textit{old}}, \text{ let } H_{\textit{old}} \coloneqq H_{\textit{new}}, \text{ calculate } \overline{X}_{H_{\textit{new}}} \text{ and let } H_{\textit{old}} \coloneqq H_{\textit{new}}, \ \overline{X}_{H_{\textit{old}}} \coloneqq \overline{X}_{H_{\textit{new}}} \text{ and } S_{H_{\textit{new}}} := S_{H_{\textit{new}}}.$ Then go to Step 3. Otherwise, the process is stopped.



Simulation Study

A simulation study is performed by generating multivariate data from the following mixture *p*-variate normal distributions [5,6,18–20] and is given as in equation (3).

$$(1-\varepsilon)N_{p}(\overset{\Gamma}{\mu_{0}},\Sigma_{0}) + \varepsilon N_{p}(\lambda \overset{\Gamma}{\mu_{1}},\delta \Sigma_{1}), \tag{3}$$

where $\Sigma_0 = \Sigma_1 = I_p$, $\overset{\Gamma}{\mu_0} = \begin{pmatrix} 0 \text{ OK } 0 \end{pmatrix}'$ and $\overset{\Gamma}{\mu_1} = \begin{pmatrix} 1 \text{ 1K } 1 \end{pmatrix}'$ is of dimension p. Inliers are generated from $N_p \begin{pmatrix} \Gamma \\ \mu_0, \Sigma_0 \end{pmatrix}$ and outliers are generated from $N_p \begin{pmatrix} \lambda \\ \mu_1, \delta \Sigma_1 \end{pmatrix}$. The separation between outliers and inliers is determined by the values of λ and δ where both values are mean shift and covariance shift values, respectively. In this study, both λ and δ are shifted simultaneously. The outlier scenario in this study is different from the study done by Abd Mutalib et al. [13], Abd Mutalib et al. [14], and Abd Mutalib et al. [15], where mean (λ) or covariance (δ) are shifted separately.

The simulation study has been conducted for different values of the sample sizes, n=30,50,100 and different number of variables, p=2,3,5. The percentage of outliers is set as $\varepsilon=5\%,15\%$, the distance of outliers by the shifting mean is $\lambda=2,4$, and the distance of outliers by shifting the covariance is $\delta=2.10$.

Performance Measurement

The steps to identify outliers are given below. Robust mean and covariance matrix obtained from FMCD, MVV, CME, ISE and TOC will replace $\overline{\mathbf{x}}$ and \mathbf{S} in Step 1.

Step 1: Compute the distance $d^2(\mathbf{x}_i) = \sqrt{(\mathbf{x}_i - \overline{\mathbf{x}})^t \mathbf{S}^{-1}(\mathbf{x}_i - \overline{\mathbf{x}})}$ for i = 1, 2, K n, where $\overline{\mathbf{x}}$ and \mathbf{S} are the robust mean and covariance matrix of FMCD, MVV, CME, ISE and TOC.

Step 2: Use the cut-off value $\sqrt{\chi_{p,0.975}^2}$ to detect outliers. If $d(\mathbf{x}_i) > \sqrt{\chi_{p,0.975}^2}$, \mathbf{x}_i is an outlier.

The simulation study was done using the R statistical package. The simulation is run repeatedly 10000 times. The performance of the five robust estimators is measured by using three different measurements, which are the probability of detecting outliers (pout), masking error (pmask) and swamping error (pswamp) given as follows [16,21]. Equations (4) – (6) are the formulas for the performance measures:

$$pout = \frac{"success"}{s},\tag{4}$$

$$pmask = \frac{\text{"failure"}}{(out)(s)},\tag{5}$$

$$pswamp = \frac{"false"}{(n - out)(s)},$$
(6)

where "success" is the number of data sets that the robust estimators successfully identified all the outliers, "failure" is the number of outliers in all data sets that are detected as inliers and "false" is the number of inliers in all data sets that falsely detected as outliers. Meanwhile, s is the total number of simulations, out is the number of outliers and s is the sample size. The s pout, s pmask and s pswamp values will be between 0 and 1. The best robust estimator will show the highest s pout value when the value approaches 1 and the lowest value of s pmask and s pswamp when the value approaches 0 [22].

Results and Discussion

Results of the simulation study are presented in Table 1 to Table 3 and illustrated in Figure 3 to Figure 5. Table 1 shows results for the probability of detecting outliers (pout) by all robust estimators. The best results are bold to indicate the best robust estimator's performance. From Table 1, the pout values of each δ for all robust estimators increase when values of λ increases for any fixed values of n, p and ε . It shows that all robust estimators have better performance detecting outliers when the values of λ increase. From Table 1, the pout values for all robust estimators are 1.0000 when p = 5, $\varepsilon = 5\%$ for all values of n. These results indicate that all robust estimators successfully detected all outliers for these



cases. Most of the *pout* values are more than 0.9000 when $\lambda=4, \varepsilon=5\%$ for any fixed values of δ, n and p. As we can see from Table 1, the *pout* values for all robust estimators are decreasing as the percentage of outliers, ε increasing except for one case when $n=30, p=2, \delta=2, \lambda=2$. This might be because as the number of outliers increases, it becomes harder for all robust estimators to detect outliers. From Table 1, TOC show quite good results and becomes the best estimator in some cases, for example, when $n=30, p=3, \delta=2, \lambda=4$ and $\varepsilon=15\%$. Alternatively, the results for *pout* can be represented in Figure 3. For illustrative purposes, only the graph for n=100 and p=5 were chosen. From Figure 3, all robust estimators approach 1 as the values of λ increasing and as we can see in this case, *pout* values approach 1 faster for $\delta=10$ than $\delta=2$.

Table 1. The performance measure using "success" probability (pout) of different robust estimators

FMCD		n	c	1			$\varepsilon = 5\%$					ε=15%		
1	n	P	0	λ	FMCD	MVV	CME	ISE	TOC	FMCD	MVV	CME	ISE	TOC
10		2	0	2	0.6636	0.6252	0.6636	0.6252	0.6636	0.4809	0.4919	0.4919	0.4809	0.4919
10 2 0.7600 0.7685 0.7630 0.7647 0.5233 0.5233 0.5233 0.9323 0.9925 0.9029 0.9029 0.9029 0.9029 0.9029 0.9029 0.9029 0.9029 0.9029 0.9029 0.9029 0.9029 0.9029 0.9085 0.6835 0.6832 0.8682 0.8682 0.8682 0.8682 0.9009 0.9981 0.9851 0.9851 0.9829 0.9829 0.9829 0.9829 0.9829 0.9829 0.9829 0.9829 0.9829 0.8830 0.8830 0.8830 0.8835 0.9835 0.9835 0.9835 0.9835 0.9835 0.9835 0.9835 0.9936 0			2	4	0.9745	0.9793	0.9793	0.9793	0.9846	0.9933	0.9940	0.9940	0.9940	0.9918
1			10	2	0.7600	0.7585	0.7535	0.7500	0.7467	0.5233	0.5233	0.5233	0.5233	0.4932
10				4	0.9029	0.9029	0.9029	0.9029	0.9029	0.6975	0.6975	0.6832	0.6975	0.6975
10	•	2	2	2	0.8682	0.8682	0.8682	0.8682	0.7776	0.2224	0.2298	0.2224	0.2224	0.2224
10 2 0.9951 0.897 0.892 0.8962 0.9009 0.7193 0.7193 0.7084 0.7193 0.7045 14 0.9704 0.9704 0.9704 0.9702 0.9699 0.8900 0.8900 0.8890 0.8895 0.8895 15 2 2 0.9582 0.9588 0.9603 0.9506 0.9592 0.8870 0.8870 0.8860 0.9996 0.8870 10 2 0.9818 0.9803 0.9806 0.9506 0.9592 0.8870 0.8860 0.9990 0.9990 1.0000 10 2 0.9818 0.9803 0.9824 0.9829 0.9822 0.9504 0.9508 0.9440 0.9542 0.9542 10 2 0.6408 0.6286 0.6600 0.	00		2	4	0.9999	0.9999	0.9999	0.9999	0.9998	0.9851	0.9851	0.9829	0.9829	0.9852
10	30	3	10	2	0.9051	0.897	0.892	0.8962	0.9009	0.7193	0.7193	0.7084	0.4809 0.9940 0.5233 0.6975 0.2224 0.9829 0.7193 0.8895 0.9096 0.9990 0.9542 0.9936 0.1625 0.9844 0.2620 0.6736 0.3538 0.9960 0.5998 0.8781 0.2384 0.9999 0.8739 0.8739 0.9872 0.0484 0.9528 0.1067 0.4375 0.0176 0.9895 0.8093	0.7045
Facility			10	4	0.9704	0.9704	0.9704	0.9702	0.9699	0.8900	0.8900	0.8900	0.8895	0.8895
Table			2	2	0.9582	0.9588	0.9603	0.9506	0.9592	0.8870	0.8870	0.8650	0.9096	0.8870
10 2 0.9818 0.9803 0.9824 0.9829 0.9822 0.9504 0.9508 0.9440 0.9542 0.9542		_	2	4	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	0.9990	1.0000
10		Page 2	0.9822	0.9504	0.9508	0.9440	0.9542	0.9542						
Parison			10	4	0.9984	0.9983	0.9986	0.9987	0.9984	0.9940	0.9936	0.9936	0.4809 0. 0.9940 0. 0.5233 0. 0.6975 0. 0.2224 0. 0.9829 0. 0.7193 0. 0.8895 0. 0.9996 0. 0.9936 0. 0.9936 0. 0.1625 0. 0.9844 0. 0.2620 0. 0.5998 0. 0.5998 0. 0.8781 0. 0.8739 0. 0.8739 0. 0.9872 0. 0.0484 0. 0.9528 0. 0.1067 0. 0.4375 0. 0.3112 0. 0.8043 0. 0.4286 0. 1.0000 1. 0.8046 0.	0.9940
10			2	2	0.6408	0.6286	0.6600	0.6600	0.6600	0.1625	0.1378	0.1378	0.1625	0.1378
Table			2	4	0.9951	0.9951	0.9951	0.9936	0.9906	0.9822	0.9822	0.9822	0.9844	0.9822
50 4 0.8382 0.8382 0.8239 0.8239 0.6736 0.8781 0.8782 0.99999 0.99999 0.99999			10	2	0.6632	0.6621	0.6632	0.6621	0.6632	0.2620	0.2620	0.2620	0.2620	0.2620
Table Tabl			10	4	0.8382	0.8382	0.8239	0.8295	0.8239	0.6736	0.6736	0.6736	0.6736	0.6437
Table Tabl			2 -	2	0.8076	0.8076	0.8076	0.6765	0.7495	0.3538	0.4212	0.3538	0.3538	0.3538
10 2 0.8250 0.8300 0.8300 0.8250 0.8300 0.5998 0.5998 0.6030 0.5998 0.5877 4 0.9554 0.9558 0.9554 0.9554 0.9558 0.8756 0.8781 0.8781 0.8781 0.8781 0.8781 5 2 2 0.9399 0.9460 0.9410 0.9460 0.9245 0.5916 0.8020 0.5913 0.2384 0.5916 4 1.0000 1.0000 1.0000 1.0000 1.0000 0.9999 0.	F O			4	0.9967	0.9967	0.9964	0.9964	0.9972	0.9961	0.9960	0.9960	0.9960	0.9933
100 1	50	3	10	2	0.8250	0.8300	0.8300	0.8250	0.8300	0.5998	0.5998	0.6030	0.9960 0.5998 0.8781	0.5877
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$			10	4	0.9554	0.9558	0.9554	0.9554	0.9558	0.8756	0.8781	0.8781		0.8781
100 4 1,0000 1,0000 1,0000 1,0000 1,0000 0,9999 0,9986 0,8726 4 0,9974 0,9980 0,9987 0,9981 0,9473 0,3181 0,0484 0,0484 0,0484 0,0484 0,0203 10 2 0,5675 0,5675 0,5705 0,5705 0,5663 0,1049 0,1067		-	2	2	0.9399	0.9460	0.9410	0.9460	0.9245	0.5916	0.8020	0.5913	0.2384	0.5916
100 2 0.9590 0.9624 0.9580 0.9593 0.9607 0.8746 0.8771 0.8683 0.8739 0.8726 4 0.9974 0.9980 0.9975 0.9981 0.9977 0.9877 0.9866 0.9865 0.9872 0.9866 2 2 0.3431 0.3641 0.3431 0.3431 0.3181 0.0484 0.0484 0.0484 0.0484 0.0203 4 0.9864 0.9864 0.9864 0.9795 0.9528 0.9528 0.9528 0.9528 0.9473 2 0.5675 0.5675 0.5705 0.5705 0.5663 0.1049 0.1067 0.1049 0.1067 0.0881 4 0.8035 0.8035 0.8035 0.8035 0.7948 0.4433 0.4375 0.4375 0.4375 0.4185 4 0.9991 0.9991 0.9991 0.9991 0.9987 0.9895 0.9895 0.9895 0.9895 100 2 0.7645 0.7630 0.7645 0.7643 0.7582 0.3112 0.3112 0.3112 0.3112 10 2 0.8326 0.8436 0.8326 0.8326 0.7972 0.4321 0.4321 0.4321 0.4286 0.3333 2 0.9317 0.9326 0.9329 0.9326 0.9317 0.7987 0.7984 0.8053 0.8046 0.7864				4	1.0000	1.0000	1.0000	1.0000	1.0000	0.9999	0.9999	0.9999	0.9999	0.9999
100 3 0.9974 0.9980 0.9975 0.9981 0.9977 0.9877 0.9866 0.9865 0.9872 0.9866 100 2 0.3431 0.3431 0.3431 0.3181 0.0484 0.0473 0.0473 0.0473 0.0473 0.0473 0.0473 0.0473 0.0473 0.0476 0.1049 0.1067 0.1049 0.1067 0.1049 0.1067 0.04375 0.4185 0.4185 0.4185 0.4185 0.4185 0.4185 0.4185 0.4185 0.0476 0.0176 0.0176 0.0176 0.0176 0.0176 0.0176 0.0176 0.0		5	10	2	0.9590	0.9624	0.9580	0.9593	0.9607	0.8746	0.8771	0.8683	0.8739	0.8726
100 2			10	4	0.9974	0.9980	0.9975	0.9981	0.9977	0.9877	0.9866	0.9865	919	0.9866
100 2			2	2	0.3431	0.3641	0.3431	0.3431	0.3181	0.0484	0.0484	0.0484	0.0484	0.0203
100 3		2		4	0.9864	0.9864	0.9864	0.9864	0.9795	0.9528	0.4919 0.4919 0.4809 0.9940 0.9940 0.9940 3 0.5233 0.5233 0.5233 5 0.6975 0.6832 0.6975 4 0.2298 0.2224 0.2224 1 0.9851 0.9829 0.9829 3 0.7193 0.7084 0.7193 0 0.8900 0.8895 0.9096 0 0.8870 0.8650 0.9096 0 0.9508 0.9440 0.9542 0 0.9936 0.9936 0.9936 0 0.9936 0.9936 0.9936 0 0.9932 0.9844 0.9542 0 0.9936 0.9936 0.9936 0 0.9822 0.9844 0.1625 0 0.9822 0.9844 0.0462 0 0.2620 0.2620 0.2620 0 0.6736 0.6736 0.6736 0 0.8781 0.8781 0.8781	0.9473		
100 3 4 0.8035 0.8035 0.8035 0.8035 0.7948 0.4433 0.4375 0.4375 0.4375 0.4375 0.4375 0.4375 0.4375 0.4375 0.4375 0.4185 100 2 0.5550 0.5550 0.5550 0.5550 0.5436 0.0176 0.0161 0.0176 0.0176 0.0161 0.0176 0.0176 0.0161 0.0161 0.0176 0.0161 0.0161 0.0176 0.0161 0.0161 0.0161 0.0161 0.0161 0.0161 0.0161 0.0161 0.0161 0.0161 0.0161 0.0161 0.0161 0.0161 0.0161 0.0161 0.0161 0.0161 0.0161 <td></td> <td>2</td> <td>10</td> <td>2</td> <td>0.5675</td> <td>0.5675</td> <td>0.5705</td> <td>0.5705</td> <td>0.5663</td> <td>0.1049</td> <td>0.1067</td> <td>1.0000 1.0000 0.9990 0.9508 0.9440 0.9542 0.9936 0.9936 0.9936 0.1378 0.1625 0.9822 0.9844 0.2620 0.2620 0.6736 0.6736 0.4212 0.3538 0.9960 0.9960 0.5998 0.6030 0.5998 0.6030 0.5998 0.6030 0.5998 0.6030 0.8781 0.8781 0.8020 0.5913 0.2384 0.9999 0.9999 0.9865 0.9872 0.0484 0.0484 0.9528 0.9528 0.9528 0.9528 0.1067 0.1049 0.1067 0.4375 0.4375 0.4375 0.3112 0.3112 0.3112 0.8093 0.8094 0.8093 0.4321 0.4286 1.0000 1.0000 0.7984 0.8053 0.8046</td> <td>0.0881</td>		2	10	2	0.5675	0.5675	0.5705	0.5705	0.5663	0.1049	0.1067	1.0000 1.0000 0.9990 0.9508 0.9440 0.9542 0.9936 0.9936 0.9936 0.1378 0.1625 0.9822 0.9844 0.2620 0.2620 0.6736 0.6736 0.4212 0.3538 0.9960 0.9960 0.5998 0.6030 0.5998 0.6030 0.5998 0.6030 0.5998 0.6030 0.8781 0.8781 0.8020 0.5913 0.2384 0.9999 0.9999 0.9865 0.9872 0.0484 0.0484 0.9528 0.9528 0.9528 0.9528 0.1067 0.1049 0.1067 0.4375 0.4375 0.4375 0.3112 0.3112 0.3112 0.8093 0.8094 0.8093 0.4321 0.4286 1.0000 1.0000 0.7984 0.8053 0.8046	0.0881	
100 3 2 4 0.9991 0.9991 0.9991 0.9991 0.9987 0.9895 0			10	4	0.8035	0.8035	0.8035	0.8035	0.7948	0.4433	0.4375	0.4375	0 0.8895 0 0.9096 0 0.9990 0 0.9542 6 0.9936 8 0.1625 2 0.9844 0 0.2620 6 0.6736 8 0.3538 0 0.9960 0 0.5998 1 0.8781 8 0.2384 9 0.9999 8 0.8739 6 0.9872 4 0.0484 8 0.9528 9 0.1067 6 0.4375 6 0.0176 6 0.9895 2 0.3112 4 0.8093 1 0.4286 0 1.0000 8 0.8046	0.4185
100 3			2	2	0.5550	0.5550	0.5550	0.5550	0.5436	0.0176	0.0176	0.0176	0.0176	0.0161
10 2 0.7645 0.7630 0.7645 0.7645 0.7582 0.3112	100	2		4	0.9991	0.9991	0.9991	0.9991	0.9987	0.9895	0.9895	0.9895	0.9895	0.9895
4 0.9381 0.9376 0.9381 0.9365 0.9310 0.8093 0.8093 0.8094 0.8093 0.8093 2 2 0.8326 0.8326 0.7972 0.4321 0.4321 0.4321 0.4286 0.3333 4 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 1.0000 0.7864	100	3	10	2	0.7645	0.7630	0.7645	0.7643	0.7582	0.3112	0.3112	0.3112		0.3112
5 4 1.0000 1.000			10	4	0.9381	0.9376	0.9381	0.9365	0.9310	0.8093	0.8093	0.8094	0.8093	0.8093
5 4 1.0000 1.000			2	2	0.8326	0.8436	0.8326	0.8326	0.7972	0.4321	0.4321	0.4321	0.4286	0.3333
2 0.9317 0.9326 0.9329 0.9326 0.9317 0.7987 0.7984 0.8053 0.8046 0.7864		5		4	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000	1.0000
10.		5	10	2	0.9317	0.9326	0.9329	0.9326	0.9317	0.7987	0.7984	0.8053	0.8046	0.7864
4 0.9944 0.9942 0.9944 0.9944 0.9944 0.9784 0.9784 0.9784 0.9793 0.9784			10	4	0.9944	0.9942	0.9944	0.9944	0.9944	0.9784	0.9784	0.9784	0.9793	0.9784



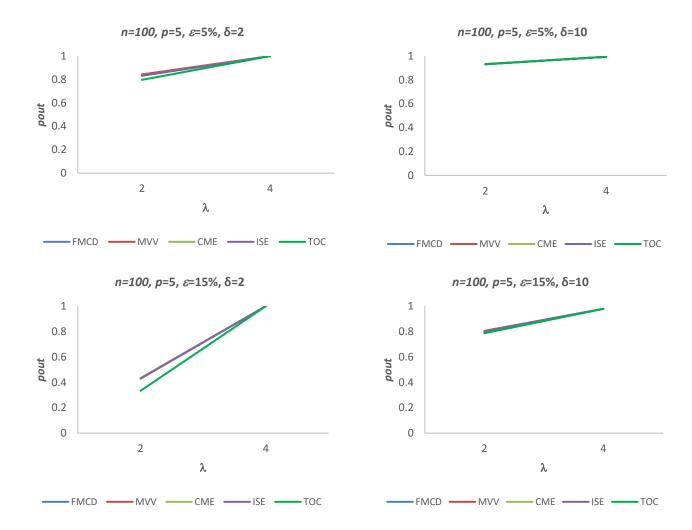


Figure 3. Plot of "success" probability (*pout*) versus distance of outliers and inliers (λ)

Table 2 shows the results for the probability of misclassifying outliers as inliers (pmask). The pmask values of each δ for all robust estimators decrease when the values of λ increases for any fixed values of n, p and ε . From Table 2, the *pmask* values for all robust estimators are 0.0000 when p=5 for any fixed values of n and ε . These results indicate that all robust estimators do not misclassify outliers as inliers for these cases. Most of the pmask values from Table 2 are less than 0.1000, which indicates all robust estimators show good performance in not misclassifying outliers as inliers in most cases. The highest pmask value is 0.2596 and is recorded by FMCD, CME, ISE and TOC when $n=30,\,p=3,\delta=2,\lambda=2$ and $\varepsilon=15\%$. The bold number is to show the best results for *pmask*. As seen from Table 2, TOC shows good results and becomes the best estimator in some cases, such as when n = 50, p = 3, $\delta = 2$, $\lambda = 4$ and $\varepsilon = 5\%$. Figure 4 shows a graphical representation for *pmask* values. We choose graph for n = 100 and p = 5 for illustrative purposes. From Figure 4, all robust estimators approach 0 as the values of λ increasing. This shows that all robust estimators have low probability of misclassifying outliers as inliers when the distance of outliers increases by shifting the mean. From Figure 4 we also can see that *pmask* values approach 0 faster as the values of λ increasing for $\delta = 10$ than $\delta = 2$. All robust estimators have lower masking errors as the values of λ and δ increasing.



Table 2. The performance measure using masking error (*pmask*) of different robust estimators

N		p	δ	2			<i>ε</i> = 5%					ε=15%		
1	n			0	λ	FMCD	MVV	CME	ISE	TOC	FMCD	MVV	CME	ISE
1			_	2	0.1868	0.2104	0.1868	0.2104	0.1868	0.1358	0.1322	0.1322	0.1358	0.1322
1		0	2	4	0.0128	0.0104	0.0104	0.0104	0.0078	0.0013	0.0012	0.0012	0.0012	0.0016
1		2	10	2	0.1280	0.1292	0.1317	0.1337	0.1358	0.1218	0.1218	0.1218	0.1218	0.1340
10 10 10 10 10 10 10 10			10		0.0496	0.0496	0.0496	0.0496	0.0496	0.0689	0.0689	0.0731	0.0689	0.0689
$ \begin{array}{l l l l l l l l l l l l l l l l l l l $			2	2	0.0679	0.0679	0.0679	0.0679	0.1181	0.2596	0.2542	0.2596	0.2596	0.2596
	00	2		4	0.0000	0.0000	0.0000	0.0000	0.0001	0.0030	0.0030	0.0034	0.0034	0.0030
	30	3	10	2	0.0489	0.0536	0.0534	0.0534	0.0511	0.0646	0.0646	0.0650	0.0646	0.0691
			10	4	0.0149	0.0149	0.0149	0.0150	0.0152	0.0230	0.0230	0.0230	0.0231	0.0231
$\begin{array}{l l l l l l l l l l l l l l l l l l l $			2	2	0.0211	0.0209	0.0200	0.0251	0.0205	0.0239	0.0239	0.0289	0.0189	0.0239
$\begin{array}{l l l l l l l l l l l l l l l l l l l $		E	2	4	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000
$\begin{array}{l l l l l l l l l l l l l l l l l l l $		Э	4.5	2	0.0092	0.0100	0.0089	0.0086	0.0090	0.0102	0.0101	0.0115	0.0094	0.0094
$\begin{array}{l l l l l l l l l l l l l l l l l l l $			10		0.0008	0.0008	0.0007	0.0006	0.0008	0.0012	0.0013	0.0013	0.0013	0.0012
$\begin{array}{l l l l l l l l l l l l l l l l l l l $			2	2	0.1376	0.1437	0.1291	0.1291	0.1291	0.2046	0.2200	0.2200	0.2046	0.2200
$\begin{array}{l l l l l l l l l l l l l l l l l l l $				4	0.0016	0.0016	0.0016	0.0021	0.0032	0.0022	0.0022	0.0022	0.0020	0.0022
$\begin{array}{l l l l l l l l l l l l l l l l l l l $			10	2	0.1273	0.1279	0.1273	0.1279	0.1273	0.1530	0.1530	0.1530	0.1530	0.1530
$\begin{array}{l l l l l l l l l l l l l l l l l l l $			10	4	0.0572	0.0572	0.0624	0.0606	0.0624	0.0479	0.0479	0.0479	0.0479	0.0536
$\begin{array}{c c c c c c c c c c c c c c c c c c c $			2	2	0.0689	0.0689	0.0689	0.1222	0.0921	0.1218	0.1024	0.1218	0.1218	0.1218
$\begin{array}{c c c c c c c c c c c c c c c c c c c $	F O			4	0.0011	0.0011	0.0012	0.0012	0.0009	0.0005	0.0005	0.0005	0.0005	0.0001
10 10 10 10 10 10 10 10	50		10	2	0.0622	0.0602	0.0602	0.0622	0.0602	0.0620	0.0620	0.0614	0.0620	0.0642
$\begin{array}{l l l l l l l l l l l l l l l l l l l $			10		0.0150	0.0149	0.0150	0.0150	0.0149	0.0166	0.0163	0.0163	0.0163	0.0163
$\begin{array}{c c c c c c c c c c c c c c c c c c c $		5		2	0.0204	0.0182	0.0200	0.0182	0.0258	0.0633	0.0273	0.0633	0.1634	0.0633
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$				4	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000
100 1				2	0.0140	0.0127	0.0143	0.0138	0.0134	0.0166	0.0161	0.0175	0.0166	0.0168
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$				4	0.0009	0.0007	0.0008	0.0006	0.0008	0.0015	0.0017	0.0017	0.0016	0.0017
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$			2	2	0.1899	0.1808	0.1899	0.1899	0.2016	0.1793	0.1793	0.1793	0.1793	0.2263
$ \begin{array}{c ccccccccccccccccccccccccccccccccccc$		2		4	0.0028	0.0028	0.0028	0.0028	0.0042	0.0032	0.0032	0.0032	0.0032	0.0036
100 4 0.0425 0.0427 0.0528 0.0528 0.0537 0.0537 0.0537 0.0537 0.0537 0.0537 0.0536 0.2356 0.2035 0.0007 0.0007 0.0007 0.0007 0.0007 0.00750 0.0750 0.0750 0.0750 0.0750 0.0750 0.0140 0.0140 0.0140 0.0140		2	10	2	0.1060	0.1060	0.1050	0.1050	0.1064	0.1432	0.1420	0.1432	0.1420	0.1523
$\begin{array}{c ccccccccccccccccccccccccccccccccccc$			10		0.0425	0.0425	0.0425	0.0425	0.0447	0.0528	0.0537	0.0537	0.0537	0.0563
100 3			2	2	0.1114	0.1114	0.1114	0.1114	0.1148	0.2356	0.2356	0.2356	0.2356	0.2408
10 2 0.0527 0.0529 0.0527 0.0526 0.0542 0.0750 0.0750 0.0750 0.0750 0.0750 0.0750 4 0.0127 0.0128 0.0127 0.0131 0.0142 0.0140 0.0140 0.0139 0.0140 0.0140 2 0.0359 0.0333 0.0359 0.0359 0.0443 0.0540 0.0540 0.0540 0.0545 0.0698 4 0.0000 0.0000 0.0000 0.0000 0.0000 0.0000 0.0000 0.0000 0.0000 0.0000 2 0.0143 0.0141 0.0141 0.0141 0.0143 0.0148 0.0148 0.0143 0.0143 0.0159	100	3		4	0.0002	0.0002	0.0002	0.0002	0.0003	0.0007	0.0007	0.0007	0.0007	0.0007
4 0.0127 0.0128 0.0127 0.0131 0.0142 0.0140 0.0140 0.0139 0.0140 0.0140 0.0140 0.0139 0.0140 0.0140 0.0140 0.0139 0.0140 0.0143	100	3	10	2	0.0527	0.0529	0.0527	0.0526	0.0542	0.0750	0.0750	0.0750	0.0750	0.0750
5 4 0.0000 0.000				4	0.0127	0.0128	0.0127	0.0131	0.0142	0.0140	0.0140	0.0139	0.0140	0.0140
5 4 0.0000 0.000			2	2	0.0359	0.0333	0.0359	0.0359	0.0443	0.0540	0.0540	0.0540	0.0545	0.0698
2 0.0143 0.0141 0.0141 0.0141 0.0143 0.0148 0.0148 0.0143 0.0143 0.0159		5		4	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000	0.0000
4 0.0011 0.0012 0.0011 0.0011 0.0015 0.0015 0.0015 0.0014 0.0015		J	10	2	0.0143	0.0141	0.0141	0.0141	0.0143	0.0148	0.0148	0.0143	0.0143	0.0159
			10	4	0.0011	0.0012	0.0011	0.0011	0.0011	0.0015	0.0015	0.0015	0.0014	0.0015



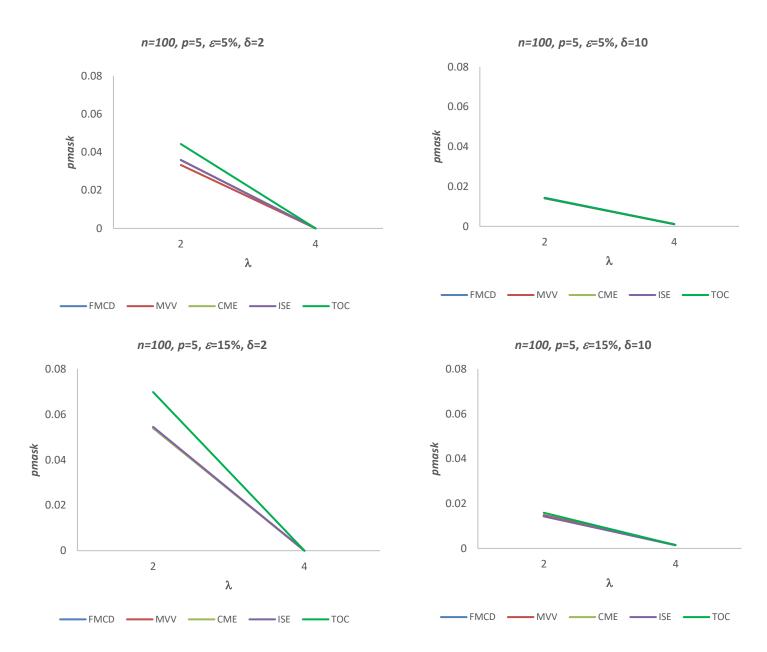


Figure 4. Plot of masking error (*pmask*) versus distance of outliers and inliers (λ)

Table 3 shows results for the probability of misclassifying inliers as outliers (pswamp). For pswamp values, there is no robust estimator obtained pswamp value 0.0000 for n and p, which indicated that all robust estimators misclassify inliers as outliers for all cases. The pswamp values also show a similar pattern as pmask values. The pswamp values of each δ for all robust estimators decrease when values of λ increases for any fixed values of n, p and ε . Overall, pswamp values for all robust estimators less than 0.3100 indicate that all robust estimators have a low probability of misclassifying inliers as outliers. Figure 5 shows all robust estimators approaching 0 as the values of λ increasing. This indicates that all robust estimators have a low probability of misclassifying inliers as outliers when the distance of outliers increases by shifting the mean.



Table 3. The performance measure using swamping error (pswamp) of different robust estimators

		δ				$\varepsilon = 5\%$					ε=15%		
n	p		λ	FMCD	MVV	CME	ISE	TOC	FMCD	MVV	CME	ISE	TOC
	2		2	0.1464	0.1211	0.1464	0.1211	0.1464	0.1817	0.1866	0.1866	0.1817	0.1866
		2	4	0.1075	0.0915	0.0915	0.0915	0.094	0.1295	0.1262	0.1262	0.1262	0.0903
		40	2	0.1755	0.219	0.1766	0.1839	0.1672	0.1713	0.1713	0.1713	0.1713	0.1320
		10	4	0.1592	0.1592	0.1592	0.1592	0.1592	0.0796	0.0796	0.0703	0.0796	0.0796
		2	2	0.2253	0.2253	0.2253	0.2253	0.1267	0.0911	0.1017	0.0911	0.0911	0.0911
00	3		4	0.1521	0.1521	0.1521	0.1521	0.114	0.0529	0.0529	0.0592	0.0592	0.0469
30	3	10	2	0.1791	0.1741	0.1790	0.1790	0.1672	0.1620	0.1620	0.1833	0.1620	0.1176
		10	4	0.1059	0.1059	0.1059	0.1390	0.0922	0.1179	0.1179	0.1179	0.0859	0.0859
		2	2	0.2705	0.3025	0.2524	0.2132	0.2356	0.1936	0.1936	0.1919	0.2029	0.1936
	_	2	4	0.2040	0.2282	0.2282	0.1985	0.2133	0.1574	0.1785	0.1883	0.1804	0.1883
	5	10	2	0.3098	0.2779	0.3060	0.2720	0.2406	0.2266	0.2347	0.2234	0.2363	0.2363
		10	4	0.3039	0.2574	0.2584	0.2431	0.2275	0.1990	0.1803	0.1803	0.1803	0.1990
		2	2	0.2463	0.2511	0.2402	0.2402	0.2402	0.1401	0.1582	0.1582	0.1401	0.1582
	2		4	0.1756	0.1756	0.1756	0.1490	0.1419	0.1222	0.1222	0.1222	0.1297	0.1222
		10	2	0.1582	0.1685	0.1582	0.1685	0.1582	0.1046	0.1046	0.1046	0.1046	0.1046
		10	4	0.1104	0.1104	0.1249	0.1315	0.1249	0.0990	0.0990	0.0990	0.0990	0.0954
	3	2	2	0.2054	0.2054	0.2054	0.2043	0.1743	0.1900	0.1849	0.1900	0.1900	0.1900
50			4	0.1375	0.1375	0.1367	0.1367	0.1172	0.1433	0.1412	0.1412	0.1412	0.1398
50		10	2	0.1634	0.1716	0.1716	0.1634	0.1716	0.1311	0.1311	0.1245	0.1311	0.1197
			4	0.1553	0.1627	0.1553	0.1553	0.1627	0.0982	0.1008	0.1008	0.1008	0.1008
	5	10	2	0.2433	0.2456	0.2370	0.2456	0.2185	0.1354	0.1346	0.1354	0.1653	0.1354
			4	0.2252	0.2393	0.2252	0.2252	0.2135	0.0925	0.0998	0.0998	0.0950	0.0925
	5		2	0.1754	0.1972	0.1772	0.1786	0.1813	0.1156	0.1155	0.1142	0.1281	0.1145
			4	0.1731	0.1655	0.1732	0.1731	0.1651	0.1030	0.0991	0.1117	0.0942	0.0991
		2	2	0.1672	0.1677	0.1672	0.1672	0.1544	0.1293	0.1293	0.1293	0.1293	0.1162
	2		4	0.1175	0.1175	0.1175	0.1175	0.1123	0.0972	0.0972	0.0972	0.0972	0.0930
	2	10	2	0.1980	0.1980	0.1917	0.1917	0.1886	0.1272	0.1305	0.1272	0.1305	0.1080
		10	4	0.1693	0.1693	0.1693	0.1693	0.1649	0.0834	0.0860	0.0860	0.0860	0.0795
		2	2	0.1144	0.1144	0.1144	0.1144	0.1101	0.0787	0.0787	0.0787	0.0787	0.0795
100	3		4	0.0956	0.0956	0.0956	0.0956	0.0947	0.0693	0.0693	0.0693	0.0693	0.0695
100	3	10	2	0.1420	0.1391	0.1420	0.1402	0.1363	0.1300	0.1300	0.1300	0.1300	0.1300
		10	4	0.1112	0.1081	0.1112	0.1094	0.1101	0.0973	0.0973	0.0985	0.0973	0.0973
		2	2	0.1370	0.1380	0.1370	0.1370	0.1359	0.1041	0.1041	0.1041	0.0998	0.0968
	5		4	0.1270	0.1261	0.1270	0.1266	0.1270	0.0886	0.0900	0.0900	0.0900	0.0832
	3	10	2	0.1286	0.1300	0.1286	0.1300	0.1286	0.1145	0.1172	0.1243	0.1205	0.1022
			4	0.1207	0.1223	0.1207	0.1207	0.1207	0.0861	0.0861	0.0861	0.0963	0.0861



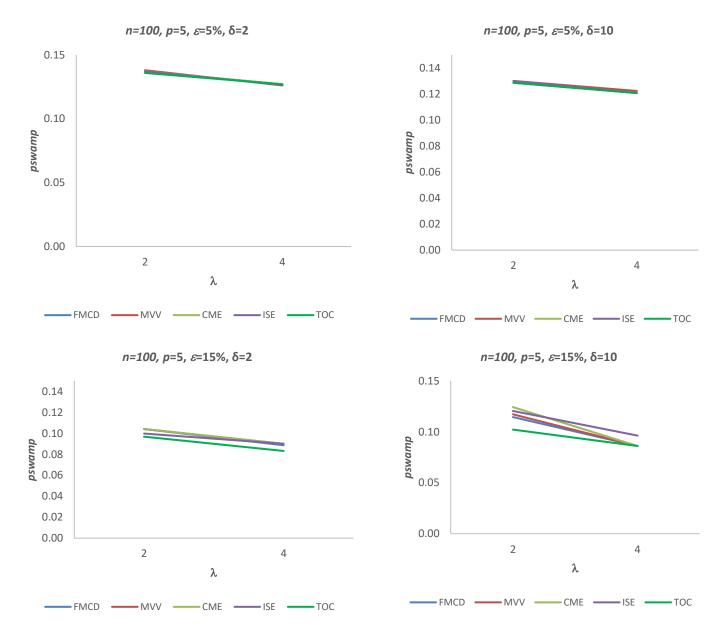


Figure 5. Plot of swamping error (pswamp) versus distance of outliers and inliers (λ)

Table 4 shows the *pout*, *pmask* and *pswamp* summary for the best robust estimator in each case. The bod entry is to show the best robust estimator. For *pout* values, it can be seen that TOC can be the best estimator or have the same performance as other robust estimators to detect outliers for most cases except when n=100, p=2. Out of 36 cases, TOC shows the best result or has a similar performance to other robust estimators in 8 cases when $\lambda=4, \varepsilon=5\%, 15\%$ and 4 cases when $\lambda=2, \varepsilon=5\%, 15\%$. This result indicates that TOC has a high probability of successfully detecting outliers when the distance between outliers and inliers by shifting the mean is large. TOC shows the best result or has similar performance to other estimators for *pmask* values in 8 cases when $\lambda=4, \varepsilon=5\%$ and 4 cases when $\lambda=2, \varepsilon=5\%$, while 9 cases when $\lambda=4, \varepsilon=15\%$ and 4 cases when $\lambda=2, \varepsilon=15\%$. This result also shows that TOC has a low probability of misclassifying outliers as inliers when the distance between outliers and inliers by shifting the mean is large. From Table 4 for *pswamp*, it can be seen that TOC is the best estimator to not misclassify inliers as outliers in most cases regardless of the values of λ and δ . TOC shows the lowest probability of swamping error compared to other robust estimators.



Table 4. Summary of the best robust estimators

					pout		pmask	pswam	p
n	p	δ	λ-	$\varepsilon = 5\%$	ε=15%	$\varepsilon = 5\%$	ε=15%	$\varepsilon = 5\%$	$\varepsilon = 15\%$
		_	2	FMCD, CME, TOC	MVV, CME, TOC	FMCD, CME, TOC	MVV, CME, TOC	MVV, ISE	FMCD, ISE
	2	2	4	TOC	MVV, CME, ISE	TOC	MVV, CME, ISE	MVV, CME, ISE	TOC
30		46	2	FMCD	FMCD, MVV, CME, ISE	FMCD	FMCD, MVV, CME, ISE	тос	тос
		10	4	ALL	FMCD, MVV, ISE, TOC	ALL	FMCD, MVV, ISE, TOC	ALL	CME
	3	2	2	FMCD, MVV, CME, ISE	MVV	FMCD, MVV, CME, ISE	MVV	тос	FMCD, CME, ISE, TOC
			4	FMCD, MVV, CME, ISE	тос	FMCD, MVV, CME, ISE	FMCD, MVV, TOC	тос	тос
	_		2	FMCD	FMCD, MVV, ISE	FMCD	FMCD, MVV, ISE	TOC	TOC
_		10	4	FMCD, MVV, CME	FMCD, MVV, CME	FMCD, MVV, CME	FMCD, MVV, CME	тос	ISE, TOC
			2	CME	ISE	CME	ISE	ISE	CME
	_	2	4	ALL	FMCD, MVV, CME, TOC	ALL	ALL	ISE	FMCD
	5 -		2	ISE	ISE, TOC	ISE	ISE, TOC	TOC	CME
		10	4	ISE	FMCD, TOC	ISE	FMCD, TOC	тос	MVV,
									CME, ISE
			2	CME, ISE, TOC	FMCD, ISE	CME, ISE, TOC	FMCD, ISE	CME, ISE, TOC	FMCD, ISE
2	2	2	4	FMCD, MVV, CME	ISE	FMCD, MVV, CME	ISE	тос	FMCD, MVV, CME, TOC
	_		2	FMCD, CME, TOC	ALL	FMCD,CME, TOC	ALL	FMCD, CME, TOC	ALL
		10	4	FMCD, MVV	FMCD, MVV, CME, ISE	FMCD, MVV	FMCD, MVV, CME, ISE	FMCD, MVV	TOC
50	3	_	2	FMCD, MVV, CME	MVV	FMCD, MVV, CME	MVV	TOC	MVV
		10	4	TOC	FMCD	TOC	TOC	TOC	TOC
			2	MVV, CME, TOC	CME	MVV, CME, TOC	CME	FMCD, ISE	TOC
			4	MVV, TOC	MVV, CME, ISE, TOC	MVV, TOC	MVV, CME, ISE, TOC	FMCD, CME, ISE	FMCD
	5 _		2	MVV, ISE	MVV	MVV, ISE	MVV	TOC	MVV
		2	4	ALL	ALL	ALL	ALL	тос	FMCD, TOC
		10	2	MVV	MVV	MVV	MVV	FMCD	CME
			4	ISE	FMCD	ISE	FMCD	TOC	ISE
		2	2	MVV	FMCD, MVV, CME, ISE	MVV	FMCD, MVV, CME ,ISE	тос	TOC
	2	_	4	FMCD, MVV, CME, ISE	FMCD, MVV, CME, ISE	FMCD, MVV, CME, ISE	FMCD, MVV, CME ,ISE	TOC	TOC
			2	CME, ISE	MVV, ISE	CME, ISE	MVV, ISE	TOC	MVV, ISE
		10	4	FMCD, MVV, CME, ISE	FMCD	FMCD, MVV, CME, ISE	FMCD	тос	FMCD
		0	2	FMCD, MVV, CME, ISE	FMCD, MVV, CME, ISE	FMCD, MVV, CME, ISE	FMCD, MVV, CME, ISE	тос	FMCD, MVV, CME, ISE
100	3 _	2	4	FMCD, MVV, CME, ISE	ALL	FMCD, MVV, CME, ISE	ALL	тос	FMCD, MVV, CME, ISE
			2	FMCD, CME	ALL	ISE	ALL	TOC	ALL
		10	4	FMCD, CME	CME	FMCD, CME	CME	MVV	FMCD, MVV, ISE, TOC
		2	2	MVV	FMCD, MVV, CME	MVV	FMCD, MVV, CME	тос	TOC
		_	4	ALL	ALL	ALL	ALL	MVV	TOC
	5		2	CME	CME	MVV, CME, ISE	CME, ISE	FMCD, CME, TOC	TOC
		10	4	FMCD, CME, ISE,	ISE	FMCD, CME, ISE,	ISE	FMCD,CME, ISE,	FMCD, MVV,



Conclusions

This study investigates the performance of TOC to detect outliers for multivariate data. The performance of TOC is compared with other robust estimators, FMCD, MVV, CME and ISE, via a simulation study. This study uses one outlier scenario by simultaneously shifting the mean and covariance with various conditions, including the number of variables, sample size and percentage of outliers. From the simulation study, TOC shows good results and a promising approach as a robust estimator to detect outliers. It has a low probability of misclassifying outliers as inliers (masking error) in multivariate data ranges from 0.000 to 0.0689. TOC especially shows good results when the distance between outliers and inliers by shifting the mean is far. TOC is the best robust estimator for swamping error for most cases where the *pswamp* ranges from 0.0469 to 0.2363, whether the mean and covariance shift values are high or low. This range is the lowest *pswamp* values compared to other robust estimators. In conclusion, TOC is an applicable and promising approach for outlier detection in multivariate data and can be incorporated with other multivariate analyses [23]. TOC can also be applied to financial or health data as long as the data is continuous and low dimensions.

Conflicts of Interest

The author(s) declare(s) that there is no conflict of interest regarding the publication of this paper.

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